

GMM DIAGNOSTICS AND SPECIFICATION CALCULATIONS

Based on AP Chap. 11

- NESTED MODELS
- TESTING MOMMENTS
- ESTIMATING ON ONE GROUP OF MOMENTS, TESTING ON ANOTHER
- NON-NESTED MODELS

NESTED MODELS

EXAMPLE: CHEN, ROLL & ROSS (1986) TEST
WHETHER “MACROECONOMIC FACTORS”
DRIVE OUT THE MARKET RETURN

$$m = b_1' f_1 + b_2' f_2$$

GIVEN FACTORS f_1 , DO WE NEED $f_2 \Rightarrow b_2 = 0$

OTHER EXAMPLE: MANAGED PORTFOLIO
SUPERIOR TO PASSIVE BUY AND HOLD?
MEAN-VARIANCE CONSTRUCTION

$$m = a + b'R^w + b'_p R^P$$

R^P : MANAGED PORTFOLIO

$$H_0 : b_p = 0$$

FINAL ILLUSTRATIVE EXAMPLE: EPSTEIN-ZIN
PREFERENCES AND CRRA PREFERENCES

Two Testing procedures

- 1. WALD TEST

$$\hat{\mathbf{b}}_2' \text{var}(\hat{\mathbf{b}}_2)^{-1} \hat{\mathbf{b}}_2 \sim \chi^2_{\dim \mathbf{b}_2}$$

- 2. LR - TYPE TEST:

$$\begin{aligned} & \text{TJ}_T(\text{RESTRICTED}) - \text{TJ}_T(\text{UNRESTRICTED}) \\ & \sim \chi^2 (\#\text{RESTRICTIONS}) \end{aligned}$$

TESTING MOMENTS

$$g_T(b) = \left\{ \begin{array}{l} g_{IT}(b) \\ \vdots \\ g_{KT}(b) \\ \vdots \\ g_{NT}(b) \end{array} \right\} g_{IT}(b)$$

ADDITIONAL SET OF MOMENT
CONDITIONS

FOR INSTANCE PRICING ERRORS ON SMALL
CAPS

$$H_0 : Eg(b) = 0$$

$$H_A : Eg_1(b) = 0, Eg(b) \neq 0$$

TEST STATISTIC

$$T\mathbf{g}_T(\hat{\mathbf{b}})'S^{-1}\mathbf{g}_T(\hat{\mathbf{b}}) - T\mathbf{g}_{IT}(\hat{\mathbf{b}}_1)'S_1^{-1}\mathbf{g}_{IT}(\hat{\mathbf{b}}_1)$$

$\sim \chi^2$ (# ELIMINATED MOMENT CONDITIONS)

NON-NESTED MODELS

MOST SDF MODELS ARE NON-NESTED

FORMAL GMM TESTS FOR NON-NESTED MODELS ARE VERY DIFFICULT TO IMPLEMENT, SEE SINGLETON (1985) AND GHYSELS AND HALL (1990)

PRESPECIFIED WEIGHTING MATRICES

INSTEAD OF (STATISTICALLY) OPTIMAL
WEIGHTING MATRIX USE PRESPECIFIED
WEIGHTING MATRIX

- ECONOMICALLY INTERESTING MOMENTS
- LEVEL PLAYING FIELD

OPTIMAL S MATRIX

CHANGES AS THE MODEL AND AS ITS
PARAMETERS CHANGE

- COMPARISON OF MODELS NOT POSSIBLE VIA J-STATISTICS BECAUSE S MATRIX NOT THE SAME

NOTE: ONE CANNOT CLAIM BETTER FIT BECAUSE SMALLER J STATISTIC

- DIFFERENT WEIGHTING MATRIX
- DIFFERENT MOMENT CONDITIONS

⇒ USE OF WEIGHTING MATRIX THAT IS INVARIANT TO MODEL & PARAMETERS

- FOREGOING EFFICIENCY LIKE OLS VS GLS
- NEAR-SINGULAR S
S IS OFTEN NEAR-SINGULAR BECAUSE ASSET RETURNS ARE HIGHLY CORRELATED
ILLUSTRATIVE EXAMPLE

$$S = \begin{pmatrix} 1 & \rho \\ \rho & 1 \end{pmatrix}$$

$$S^{-1} = \frac{1}{1-\rho^2} \begin{pmatrix} 1 & -\rho \\ -\rho & 1 \end{pmatrix}$$

CHOLESKI FACTORIZATION

$$C'C = S^{-1}$$

$$C = \begin{pmatrix} (1-\rho^2)^{-1/2} & -\rho(1-\rho^2)^{-1/2} \\ 0 & 1 \end{pmatrix}$$

THEN

$$\text{Min } \mathbf{g}'_T \mathbf{S}^{-1} \mathbf{g}_T$$

$$\text{Min } (\mathbf{g}'_T \mathbf{C}')(\mathbf{C} \mathbf{g}_T)$$

$\mathbf{C} \mathbf{g}_T$ AS $\rho \rightarrow 1$ THEN FIRST ROW OF \mathbf{C} BECOMES DOMINANT WITH TWO ELEMENTS WHICH ARE LARGE AND OF OPPOSITE SIGN

SOME PRESPECIFIED WEIGHTING MATRICES

HANSEN AND JAGANATHAN (1992)

ADVOCATE USING

$$W = E(xx')^{-1}$$

SECOND MOMENT MATRIX OF PAYOFFS

ECONOMICALLY INTERESTING DISTANCE
MEASURE BETWEEN A “CANDIDATE
DISCOUNT FACTOR” AND SPACE OF TRUE
DISCOUNT FACTORS

DISTANCE BETWEEN y (MODEL FOR m) AND NEAREST VALID m IS SAME DISTANCE BETWEEN $\text{PROJ}(y|\underline{X})$ AND x^*

RECALL \underline{X} IS SPACE OF ALL AVAILABLE PAYOFFS

$$\begin{aligned}x^* &= \text{Proj}(m|\underline{X}) \\ &= p'(E_{xx'})^{-1}x\end{aligned}$$

$$\text{Proj}(y|\underline{X}) = E(yx')(E_{xx'})^{-1}x$$

RECALL ALSO $\exists! x^* \in \underline{X}$:

$$p = E(x^*x) \quad \forall x \in \underline{X}$$

$$\|\text{Proj}(y|\underline{X}) - x^*\| =$$

$$\|E(yx') E(xx')^{-1}x - p' E(xx')^{-1}x\| =$$

$$\|(E(yx') - p') E(xx')^{-1}x\| =$$

$$(E(yx - p))' E(xx')^{-1} xx' E(xx')^{-1} x (E(yx - p)) =$$

$$(E(yx - p))' (E(xx')^{-1} (E(yx - p))) = g_T' E(xx')^{-1} g_T$$

SO MINIMIZING EUCLIDEAN DISTANCE BETWEEN x^* AND $\text{Proj}(y|\underline{X})$ IS EQUIVALENT TO MINIMIZING GMM CRITERION WITH MOMENT CONDITIONS $Eyx - p = 0$ AND $W = E(xx')^{-1}$

BECAUSE THE WEIGHTING MATRIX REMAINS INVARIANT, WE CAN COMPARE ASSET PRICING MODELS TO SEE HOW CLOSE THEY COME TO x^*

$$\|\text{Proj}(y_1|\underline{X}) - x^*\|$$

$$\|\text{Proj}(y_2|\underline{X}) - x^*\|$$

IMPORTANT NOTE: ONE HAS TO KEEP THE ASSET BASE X FIXED!

IDENTITY MATRIX

$E(xx')$ IS ALSO OFTEN NEAR-SINGULAR WITH MANY ASSETS.

IDENTITY MATRIX AVOIDS (MOST) OF THESE PROBLEMS

ADVANTAGE OF $E_{xx'}$ IS THAT IT IS INVARIANT TO PORTFOLIO FORMATION

TAKE Ax INSTEAD OF x FOR NONSINGULAR MATRIX

A (NO INFORMATION IS LOST)

$$\begin{aligned} [E(yAx) - (Ap)]' E(Axx' A')^{-1} x [E(yAx) - Ap] &= \\ [E(yx) - p]' A' (A')^{-1} E(xx')^{-1} A^{-1} x A [Eyx - p] &= \\ [E(yx) - p]' E(xx')^{-1} [Eyx - p] \end{aligned}$$

PROPERTY ALSO HOLDS FOR OPTIMAL S^{-1}

Hansen-Jagannathan Bounds

Based on AP Chap. 11 and CLM Chap 8

- How to compare various models without observing the SDF.
- Recall the equations:

$$1 = E_t m_{t+1} R_{t+1}$$

$$1 = \left(E_t m_{t+1} \right) \left(E_t R_{t+1} \right) + \text{Cov}_t \left(m_{t+1}, R_{t+1} \right)$$

Let us fix the mean of $E m_{t+1} = M$ and let us consider the implied SDF $m_{t+1}(M)$ which we can write as:

$$m_{t+1}(M) = M + \left[R_{t+1} - E(R_{t+1}) \right] \beta_M$$

$$\text{Therefore } \iota = ME \left(R_{t+1} \right) + Cov \left(R_{t+1}, m_{t+1}(M) \right)$$

$$\iota = ME \left(R_{t+1} \right) + E \left(R_{t+1} - ER_{t+1} \right) \left(m_{t+1}(M) - M \right)$$

$$\iota = ME \left(R_{t+1} \right) + E \left(R_{t+1} - ER_{t+1} \right) \left(R_{t+1} - ER_{t+1} \right) \beta_M$$

$$\Rightarrow \beta_M = \Omega^{-1} \left(\iota - ME(R_{t+1}) \right)$$

$$\begin{aligned}
\text{Var}\left(m_{t+1}(M)\right) &= \beta_M' \Omega \beta_M \\
&= \left[\Omega^{-1} \left(I - \text{ME}(R_{t+1}) \right) \right]' \Omega \left[\Omega^{-1} \left(I - \text{ME}(R_{t+1}) \right) \right] \\
\text{Var}\left(m_{t+1}(M)\right) &= \left(I - \text{ME}(R_{t+1}) \right)' \Omega^{-1} \left(I - \text{ME}(R_{t+1}) \right)
\end{aligned}$$