

Adaptive Wavelet Decompositions of Stationary (Gaussian) Processes

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General Remarks

References

- [DP06a] "Gaussian stationary processes: adaptive wavelet decompositions, discrete approximations and their convergence", G. Didier and V. Pipiras, 2006, Preprint.
- [DP06b] "Adaptive wavelet decompositions of stationary time series", G. Didier and V. Pipiras, 2006, Preprint.

About

Particular, called adaptive, wavelet decompositions (AWD), of **general** stationary (Gaussian) processes X in **continuous time** [DP06a] and **discrete time** [DP06b]. They are characterized by:

- **independent** (uncorrelated) detail coefficients $d_{j,n}$;
- possibly **correlated** approximations coefficients $X_{j,n}$;
- **non(bi)orthogonal** wavelet basis (adapted to the covariance of X);
- **FWT-like** algorithm (with filters possibly depending on scale j).

In addition,

- [DP06a]: $2^{j/2}X_{j,[2^j t]} \approx X(t)$ at small scales ("wavelet crime"). Practical conditions for this to take place are provided;
- [DP06b]: associated low and high pass filters **decay to zero fast**. Zero moments play here a key role.

Other issues addressed

- Riesz bases [DP06a].
- Convergence of discrete approximations $X_{j,n}$ [DP06a].
- Many examples [DP06a,b].

On significance of AWD

- New wavelet-based decompositions with **independent** coefficients.
- It works for **general** stationary processes; AWD in continuous time was previously obtained only for fractional Brownian motion.
- It allows for practical applications: **simulations** in [DP06a,b], **MLE** in [DP06b], etc.

Other references

- V. Ahn et al. "Fractional-order regularization and wavelet approximation to the inverse estimation problem for random fields", Journal of Multivariate Analysis, 2003.
- A. Benassi, and S. Jaffard, "Wavelet decomposition of one- and several-dimensional Gaussian processes", Recent Advances in Wavelet Analysis, 1994.
- D. Donoho, "Nonlinear solution of linear inverse problems by wavelet-vaguelette decomposition", Applied and Computational Harmonic Analysis, 1995.
- Y. Meyer, F. Sellan, and M. Taqqu, "Wavelets, generalized white noise and fractional integration: the synthesis of fractional Brownian motion", The Journal of Fourier Analysis and Applications, 1999.
- J. Zhang, and G. Walter, "A wavelet-based KL-like expansion for wide-sense stationary random processes", IEEE Transactions on Signal Processing, 1994, [ZW94].

[DP06a] Example: Ornstein-Uhlenbeck Process

Representation and earlier decomposition

It is known that

$$X(t) = \int_{-\infty}^t e^{-\lambda(t-u)} dB(u) = \int_{-\infty}^{\infty} e^{itx} \frac{1}{\lambda + ix} d\widehat{B}(x) \\ \stackrel{[ZW94]}{=} \sum_n a_{J,n} \theta^J(t - 2^{-J}n) + \sum_{j \geq J} \sum_n d_{j,n} \Psi^j(t - 2^{-j}n),$$

where $a_{J,n}, d_{j,n}$ are **independent** $N(0, 1)$ and (with a scaling function ϕ and a wavelet ψ of an orthogonal MRA)

$$\widehat{\theta}^J(x) = (\lambda + ix)^{-1} 2^{-J/2} \widehat{\phi}(2^{-J}x), \quad \widehat{\Psi}^j(x) = (\lambda + ix)^{-1} 2^{-j/2} \widehat{\psi}(2^{-j}x).$$

AWD

$$X(t) \stackrel{[DP06a]}{=} \sum_n X_{J,n} \Phi^J(t - 2^{-J}n) + \sum_{j \geq J} \sum_n d_{j,n} \Psi^j(t - 2^{-j}n),$$

where X_J is AR(1) time series with

$$\widehat{X}_J(x) = 2^{-J} (1 - e^{-(\lambda 2^{-J} + ix)})^{-1} \widehat{a}_J(x), \quad \widehat{\Phi}^J(x) = \frac{2^{J/2} (1 - e^{-2^{-J}(\lambda + ix)})}{\lambda + ix} \widehat{\phi}(2^{-J}x).$$

FWT-like algorithm (at reconstruction)

$$X_{j+1} = u_j * \uparrow_2 X_j + v_j * \uparrow_2 d_j,$$

where (with Conjugate Mirror Filters u, v)

$$\widehat{u}_j(x) = \frac{1}{2} (1 + e^{-(\lambda 2^{-j} + ix)}) \widehat{u}(x), \quad \widehat{v}_j(x) = 2^{-(j+1)} (1 - e^{-(\lambda 2^{-j} + ix)})^{-1} \widehat{v}(x).$$

Because of zero moments, v_j becomes essentially finite at small scales as $j \rightarrow \infty$ (if v is finite).

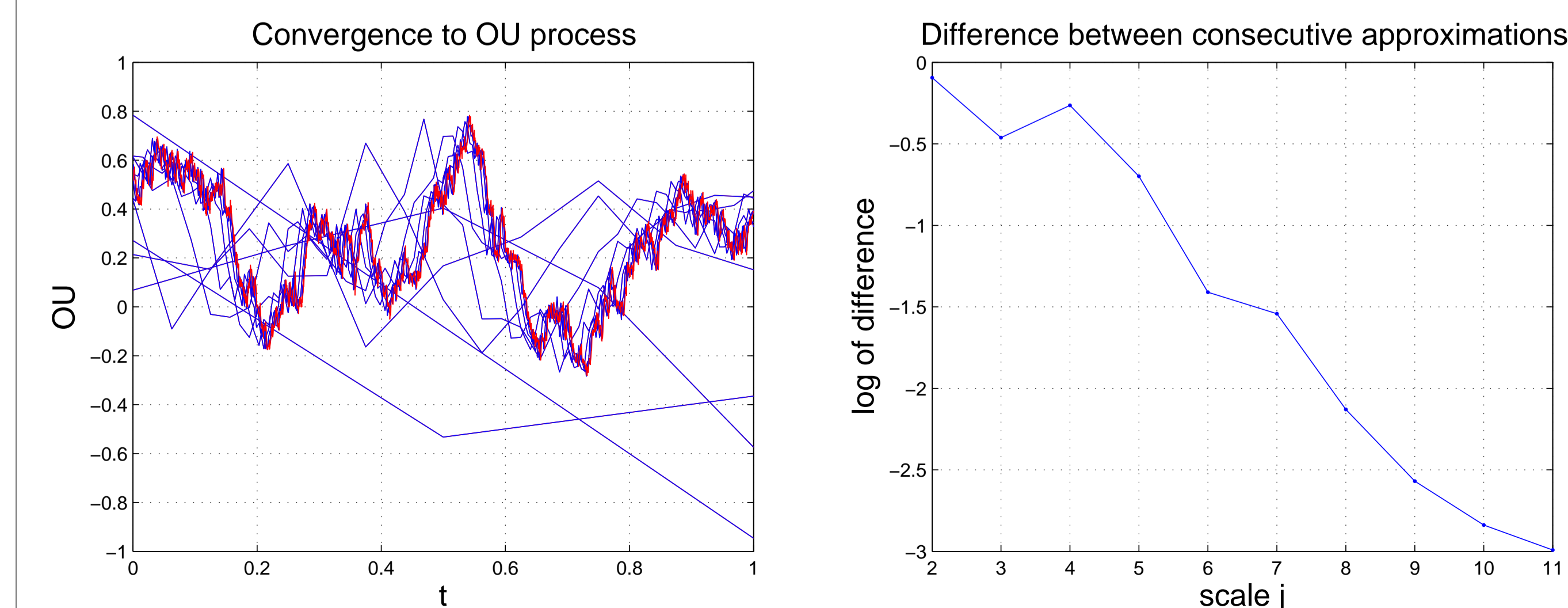
"Wavelet crime"

$$\sup_{t \in K} |2^{J/2} X_{J,[2^j t]} - X(t)| \leq C 2^{-J\gamma} \quad \text{a.s.},$$

where K is a compact interval and $\gamma > 0$ is a sample path smoothness parameter. The key here is that

$$2^J \frac{(1 - e^{-2^{-J}(\lambda + ix)})}{\lambda + ix} \approx 1, \quad \text{as } J \rightarrow \infty.$$

Simulation based on AWD



[DP06b] Example: FARIMA(0, s, 0) time series

Representation

$$X_0 = (I - B)^{-s} \epsilon^0, \quad s \in (-1/2, 1/2)$$

AWD

$$X_j = \downarrow_2 (\overline{U} * X_{j-1}), \quad d_j = \downarrow_2 (\overline{V} * X_{j-1}), \quad j \geq 1,$$

with

$$\widehat{U}(\omega) = (1 + e^{i\omega})^{-s} \widehat{u}(\omega), \quad \widehat{V}(\omega) = (1 - e^{i\omega})^s \widehat{v}(\omega).$$

One can show that d_j are **independent white noise sequences**, and X_j are **FARIMA(0, s, 0)** time series.

Filter decay vs zero moments

$(1 + e^{i\omega})^{-s}, (1 - e^{i\omega})^s$ decay extremely slowly, but not $\widehat{U}(\omega), \widehat{V}(\omega)$, since, for instance,

$$\widehat{U}(\omega) = (1 + e^{i\omega})^{-s+N} \widehat{u}_{0,N}(\omega),$$

where N is the number of zero moments of the orthogonal MRA.

In practice

Replace $*$ by \otimes . Fast decay of U, V implies **smaller border effect**.

Gaussian MLE based on AWD

For a data row vector X^0 , the negative log-likelihood is

$$\log |\Sigma_\theta| + X^0 \Sigma_\theta^{-1} X^{0'} \approx \frac{T}{\pi} \int_0^{2\pi} \log |\widehat{a}_\theta(\omega)| d\omega + Y_\theta Y_\theta',$$

where θ indicates unknown parameters, T is the size of X^0 ,

$$Y_\theta = X^0 M_\theta$$

is the detail vector in AWD with \otimes , and M_θ is an invertible AWD matrix. The key to have above is the approximate covariance matrix factorization

$$\Sigma_\theta^{-1} \approx M_\theta M_\theta'.$$

Potential advantages of MLE based on AWD

Advantages over MLE based on orthogonal wavelet decompositions: the exact MLE based on orth. wavelet decompositions involves **complex dependence structure** of its detail coefficients. By construction, detail coefficients are independent, $N(0, 1)$ variables in AWD.

Advantages over non-wavelet based MLE: MLE based on AWD is potentially **invariant to polynomial trends** in data.

MLE results for FARIMA(0, s, 0)

True parameter	Sample size T	Whittle		N zero moments	AWD		Orth. WD	
		bias	rMSE		bias	rMSE	bias	rMSE
$s = 0.3$	2^{10}	0.0003	0.0246	2	-0.0069	0.0270	-0.0079	0.0256
				6	-0.0055	0.0258	-0.0082	0.0253
	2^{14}	0.0002	0.0062	2	-0.0028	0.0068	-0.0026	0.0065
				6	-0.0027	0.0067	-0.0028	0.0068